# Example 2: Empty Models and Level-2 Predictors in General Multilevel Models for Two-Level Nested Outcomes (complete syntax, data, and output available for STATA, R, and SAS electronically)

The data for this example come from the High School and Beyond dataset (HSB4) used in McNeish (2023). Using 7,185 students from 160 schools, we will be examining the extent to which student math can be predicted from school-level variables of school size and public versus private. Note that this example computes total- $R^2$  and pseudo- $R^2$  in SAS using two custom macros (available in the SAS syntax file online) as well as in R using two custom functions and a general package (available in the R syntax file and function files online).

### **<u>STATA</u>** Syntax for Importing and Preparing Data for Analysis:

```
// Define global variable for file location to be replaced in code below
```

- // \\Client\ precedes path in Virtual Desktop outside H drive
- global filesave "C:\Dropbox\23\_PSQF6272\PSQF6272\_Example2"
- // Open trimmed example excel data file from sheet "HSB4" and clear away any existing data clear // clear memory in case of open data import excel "\$filesave\Example2 Data.xlsx", firstrow case(preserve) sheet("HSB4") clear

```
display "STATA Descriptive Statistics for Example 2 Variables" summarize math size private
```

Variable	Label	N	Mean	Std Dev	Minimum	Maximum
math	math	7185	12.7478526	6.8782457	-2.8320000	24.9930000
size	size	7185	1056.86	604.1724993	100.0000000	2713.00
private	private	7185	0.4931106	0.4999873	0	1.0000000

```
// Center and re-scale school size
  gen size100 = (size-1000)/100 // private is already 0/1
// =:!!
```

```
// Filter to only cases complete on all variables to be used below
  egen nmiss=rowmiss(math size private)
  drop if nmiss>0
```

<u>R</u> Syntax for Importing and Preparing Data for Analysis (after loading packages *readxl*, *TeachingDemos*, *psych*, *lme4*, *lmerTest*, *performance*, *nlme*, *psychometric*, and *r2mlm*):

```
# Define variables for working directory and data name -- CHANGE THESE
filesave = "C:\\Dropbox/23 PSQF6272/PSQF6272 Example2/"
filename = "Example2 Data.xlsx"
setwd(dir=filesave)
# Load Jonathan's custom R functions from folder within working directory
functions = paste0("R functions/",dir("R functions/"))
temp = lapply(X=functions, FUN=source)
# Import trimmed example excel data file in sheet "HSB4"
Example2 = read_excel(paste0(filesave,filename), sheet="HSB4")
# Convert to data frame to use in analysis
Example2 = as.data.frame(Example2)
print("R Descriptive Statistics for Example 2 Variables")
describe(x=Example2[ , c("math","size","private")])
# Center and re-scale school size
Example2$size100 = (Example2$size-1000)/100 # private is already 0/1
# Filter to only cases complete on all variables to be used below
Example2 = Example2[complete.cases(Example2[ , c("math","size","private")]),]
```

### Model 0: Single-Level Empty Means, NO Random Intercept (2 parameters)

 $Math_{pc} = \beta_0 + e_{pc}$ 

The **||** schoolID: , indicates the level-2 nesting variable, where any random effects would go after the colon. The **noconstant** option removes the default random intercept variance (for now).

#### **<u>STATA</u>** Syntax and Output for Model 0:

display "STATA Model 0: Single-Level Empty Means, NO Random Intercept" mixed math , || schoolID: , noconstant reml dfmethod(residual) dftable(pvalue) nolog

Mixed-effects REML regressic DF method: Residual	n	Number of DF:	obs = min = avg = max =	7,185 7,184.00 7,184.00 7,184.00
Log restricted-likelihood =	-24051.459 = LL	F(0, 7184 Prob > F	= .00) =	
math   Coef.	Std. Err.	DF t	. P> t	
tons   12.74785	.0811455	7184.0 157.1	0 0.000	Beta0 (overall mean)
Random-effects Parameters	Estimate	Std. Err.	[95% Conf.	Interval]
var(Residual)	47.31026	.789327	45.78823	48.88289 <b>Var(e_pc)</b>

#### display "-2LL = " e(ll)\*-2 // Print -2LL for model -2LL = 48102.917

estat ic, n(160) // AIC and BIC for # level-2 units
Akaike's information criterion and Bayesian information criterion
Model | Obs ll(null) ll(model) df AIC BIC
. | 160 . -24051.46 2 48106.92 48113.07

Note: N=160 used in calculating BIC.

#### <u>R</u> Syntax and Output for Model 0 (using gls from nlme to omit random intercept variance for now):

```
print("R Model 0: Single-Level Empty Means, NO Random Intercept")
Model0 = gls(data=Example2, model=math~1)
print("Show results including -2LL and residual variance")
-2*logLik(Model0); summary(Model0)
'log Lik.' 48102.917 (df=2) \rightarrow -2LL for model (with 2 parameters)
Generalized least squares fit by REML
 Model: math ~ 1
  Data: Example2
       AIC
                BIC
                        logLik
  48106.917 48120.676 -24051.459
Coefficients:
                Value Std.Error t-value p-value
(Intercept) 12.747853 0.081145473 157.09875 0 Beta0 (overall mean)
Residual standard error: 6.8782457
Degrees of freedom: 7185 total; 7184 residual
```

# summary(Model0)\$sigma^2 [1] 47.310264 → Var(e pc) as residual variance

# Model 1: Two-Level Empty Means, WITH Random Intercept (3 parameters)

Level 1: $Math_{pc} = \beta_{0c} + e_{pc}$	
Level 2: $\beta_{0c} = \gamma_{00} + U_{0c}$	Denominator degrees of freedom (DDF) now uses
<b><u>STATA</u></b> Syntax and Output for Model 1:	Kenward-Roger is better for really small samples).
display "STATA Model 1: Two-Level Empty Mea mixed math ,    schoolID: , reml dfmethod(:	ans, WITH Random Intercept" satterthwaite) dftable(pvalue) nolog
Mixed-effects REML regression	Number of obs = 7,185
Group variable: schoolID	Number of groups = 160
	Obs per group:
	$\min = 14$
	avg = 44.9
DE method. Cettenthusite	max = 6/
DF method: Satterthwalle	DF: $MIR = 158.82$
	avy = 150.02
	F(0, 0, 00) =
Log restricted-likelihood = -23558.397 = Li	$\mathbf{L}  \text{Prob} > F \qquad = \qquad .$
math   Coef. Std. Err.	DF t P> t
	158.8 51.71 0.000 gamma00
The fixed intercept is now the grand mean of the school r the model without a random intercept variance) wheneve	neans, which will differ from the overall grand mean (as given by r level-2 units have different level-1 sizes (i.e., are "unbalanced").

\_\_\_\_\_ Random-effects Parameters | Estimate Std. Err. [95% Conf. Interval] -----+ schoolID: Identity var(\_cons) | 8.614081 1.078813 6.739162 11.01062 Var(U\_0c) \_\_\_\_\_ var(Residual) | **39.14832** .6606445 37.87466 40.46481 **Var(e\_pc)** \_\_\_\_\_ LR test vs. linear model: chibar2(01) = 986.12 Prob >= chibar2 = 0.0000

The chibar2 test above is a likelihood ratio (LR) test comparing this model to a single-level regression (without a random intercept, as linear model) using a chi-square ( $\chi^2$ ) distribution with a mixture of DF=0 (for which  $\chi^2 = 0$  always) and DF=1. Consequently, in this case you can obtain the mixture *p*-value by weighting each contribution to the  $\chi^2$  by 0.5, which means cutting the regular *p*-value in half. Here, this LRT is a significance test of the intraclass correlation (ICC), which in turn provides an effect size for the amount of constant dependency attributed to school mean differences in math.

display "-2LL = " e(ll)*-2 -2LL = 47116.793	// Pr	int -2LL for	model	Likelihood R = 48,102.91	atio Test (LRT 7 — 47,116.79	') Statistic: 93 = 986.12
<b>estat ic, n(160)</b> Akaike's information criterion	// AI and Ba	<b>C and BIC fo</b> yesian infor	<b>r # leve</b> mation o	<b>el-2 units</b> criterion		
Model   Obs ll(	null)	ll(model)	df	AIC	BIC	
.   160	·	-23558.4	3	47122.79	47132.02	
Note: N=160 used in calculatin estat icc // Intraclass corre Intraclass correlation	g BIC. lation	$ICC = \frac{\tau}{\tau_{U_0}^2}$	$\frac{\sigma_0^2}{1+\sigma_e^2} = \frac{1}{8}$	8.614 .614 + 39.148	=. 180	
Level		ICC Std.	 Err.	[95% Conf.	Interval]	
schoolID	.180	<b>3528</b> .0187	219	.1465168	.2199886	

display "ICC2 = " 8.614081/(8.614081+(39.14832/45))
ICC2 = .90827091 → Reliability of the school means using mean cluster size

<u>R</u> Syntax and Output for Model 1 (using lmer from lme4 instead):

```
print("R Model 1: Two-Level Empty Means, WITH Random Intercept")
Model1 = lmer(data=Example2, REML=TRUE, formula=math~1+(1|schoolID))
print("Show results using Satterthwaite DDF including -2LL as deviance")
llikAIC(Model1, chkREML=FALSE); summary(Model1, ddf="Satterthwaite")
'log Lik.' -23558.397 (df=3) → LL for model (with 3 parameters)
$AICtab
       AIC
                  BIC
                          loqLik
                                  deviance
                                               df.resid
 47122.793 47143.433 -23558.397 47116.793
                                               7182.000 \rightarrow Deviance = -2LL for model
Random effects:
 Groups Name Variance Std.Dev.
 schoolID (Intercept) 8.614 2.9350
                                         Var(U_pc)
 Residual 39.148 6.2569
                                         Var(e_pc)
Number of obs: 7185, groups: schoolID, 160
Fixed effects:
             Estimate Std. Error
                                         df t value Pr(>|t|)
(Intercept) 12.63697 0.24439 156.64732 51.708 < 2.2e-16 gamma00 = mean of school means
print("Show intraclass correlation and its LRT")
icc(Model1); ranova(Model1)
# Intraclass Correlation Coefficient
                                        ICC = \frac{\tau_{U_0}^2}{\tau_{U_0}^2 + \sigma_e^2} = \frac{8.614}{8.614 + 39.148} = .180
    Adjusted ICC: 0.180
  Unadjusted ICC: 0.180
```

ANOVA-like ta	able for	random-e	effects:	Single	terr	n deletions	
	npar	logLik	AIC	LRT	Df	Pr(>Chisq)	Likelihood Ratio Test (LRT) Statistic:
<none></none>	3	-23558.4	47122.8				-2(-235584+240515) - 986124
(1   schoolII	D) 2	-24051.5	48106.9	986.124	1	< 2.22e-16	= 2(25,550.4 + 24,051.5) = 500.124

The LRT above is a likelihood ratio test comparing this model to a single-level regression (without a random intercept) using a chi-square ( $\chi^2$ ) distribution with a regular DF=1 distribution, instead of a mixture of DF=0 (for which  $\chi^2 = 0$  always) and DF=1 as the default in STATA MIXED. Consequently, in this case you can obtain the mixture *p*-value by weighting each contribution to the  $\chi^2$  by 0.5, which means cutting the regular *p*-value in half. Here, this LRT is a significance test of the intraclass correlation (ICC), which in turn provides an effect size for the amount of constant dependency attributed to school mean differences in math.

```
print("Show ICC2 for Reliability of School Mean Math -- Weighted and Unweighted")
ICC2.lme(data=Example2, dv=math, grp=schoolID, weighted=TRUE) [1] 0.90877647
ICC2.lme(data=Example2, dv=math, grp=schoolID, weighted=FALSE) [1] 0.9013773
```

**Design effect** using mean #students per school: =  $1 + ((n-1) * ICC) \rightarrow 1 + [(45-1)*.180] = 7.357$ 

Effective sample size: Effective N = (#Total Obs) / Design Effect  $\rightarrow$  7,185 / 7.357 = 977!!!

This means that our power to detect effects of level-1 person predictors will be approximately that of an independent sample of 977 students—only if the ICC were 0 would we have power for level-1 effects based on the actual number of students. Power for level-2 effects is based on the number of schools (160).

Random intercept 95% confidence interval:  $CI = \gamma_{00} \pm z_{crit} \sqrt{\tau_{U_0}^2} = 12.637 \pm 1.96\sqrt{8.614} = 6.884 \ to \ 18.390$ 

This means that 95% of the schools are expected to have school mean math outcomes between 6.884 and 18.390 (around the average of 12.637).

# Model 2: Add Main Effects of Level-2 School Size and Public vs Private (5 parameters)

Level 1:  $Math_{pc} = \beta_{0c} + e_{pc}$ Level 2:  $\beta_{0c} = \gamma_{00} + \gamma_{01}(size_c - 100) + \gamma_{02}(private_c) + U_{0c}$ 

#### **STATA** Syntax and Output for Model 2:

Mixed-effects R Group variable:	EML regressic schoolID	on	Number of obs = 7,185 Number of groups = 160 Obs per group:				
					min = avg = max =	14 44.9 67	
DF method: Satt	erthwaite		DF:		min = avg = max =	154.86 157.66 159.19	
Log restricted-	likelihood =	-23541.156	<b>F(2,</b> Prob >	157.31 F	_) = =	<b>21.80</b> 0.0000	
math	Coef.	Std. Err.	DF	t	P> t		
size100   private   cons	.0613509 3.151499 11.18194	.0389774 .4896777 .3210705	158.9 154.9 159.2	1.57 6.44 34.83	0.117 0.000 0.000	gamma01 gamma02 gamma00	
Random-effect	s Parameters	Estimate	Std. Err.	[95	5% Conf	. Interval]	
schoolID: Ident	ity var(_cons)	6.618341	.8603554	5.1	.29755	8.538894	Var(U_0c)
	var(Residual)	39.15	.6607004	37.	87623	40.4666	Var(e_pc)
LR test vs. lin	ear model: ch	nibar2(01) = 712	.02	Prob >=	= chiba:	r2 = 0.0000	
display "-2LL = " e(ll)*-2 // Print -2LL = 47082.313			LL for mode	The (con	re is still <i>iditional</i>	significant scho ICC = .144 vs .	ool dependency 180 before).

estat ic, n(160)	// AIC and	BIC for # le	evel-2 units
Nindle information on	theuten and Deveate		

Akaike's i	information	crite	rion and	Bayesian	information	criterion	
Мос	del	Obs	ll(null)	ll(mode	l) df	AIC	BIC
	·	160		-23541.	16 5	47092.31	47107.69

Note: N=160 used in calculating BIC.

predict predma corr math pred	ain Mmain		Save fixed-effect predicted outcomes Get total r to make R2
	math	predmain	
math   predmain	1.0000 0.2076	1.0000	

display "Total-R2 = " r(rho)^2 // Print total R2 relative to empty model Total-R2 = .04310206

The model F-test printed in the first part of the output, F(2, 157.31) = 21.80, p < .0001, is analogous to the F-test of the model R<sup>2</sup> in single-level regression. Here, it provides a significance test for the total-R<sup>2</sup> = .043.

#### <u>R</u> Syntax and Output for Model 2 (using lmer from lme4):

print("R Model 2: Add Main Effects of Level-2 School Size and Public vs Private") Model2 = lmer(data=Example2, REML=TRUE, formula=math~1+size100+private+(1|schoolID)) print("Show results using Satterthwaite DDF including -2LL as deviance") llikAIC(Model2, chkREML=FALSE); summary(Model2, ddf="Satterthwaite") 'log Lik.' -23541.156 (df=5) → LL for model (with 5 parameters) \$AICtab AIC BIC logLik deviance df.resid 47092.313 47126.711 -23541.156 **47082.313** 7180.000 → Deviance = -2LL for model Random effects: Groups Name Variance Std.Dev. schoolID (Intercept) 6.6183 2.5726 Var(U\_Oc) 39.1500 6.2570 Var(e pc) Residual Number of obs: 7185, groups: schoolID, 160 Fixed effects: Estimate Std. Error df t value Pr(>|t|) 0.321070 156.319744 34.8271 (Intercept) 11.181939 < 2.2e-16 gamma00 size100 0.061351 0.038977 156.069668 1.5740 0.1175 gamma01 0.489676 152.059267 6.4359 0.00000001516 gamma02 private 3.151499 Intercept  $\gamma_{00} =$ size100  $\gamma_{01} =$ private  $\gamma_{02} =$ print("F-Test of Model Total-R2") contestMD(Model2, ddf="Satterthwaite", L=rbind(c(0,1,0),c(0,0,1))) Mean Sq NumDF DenDF F value Sum Sq Pr(>F)1 1707.2093 853.60464 2 154.47003 **21.803438** 0.000000045647872 print("Psuedo-R2 relative to empty model using Jonathan's function") pseudoRSquaredinator(smallerModel=Model1, largerModel=Model2) R2 Random. (Intercept) R2 L1.sigma2 Pseudo-R<sup>2</sup> Results: The two fixed slopes of our two 0.231683234632 -0.000042893403 level-2 predictors accounted for 23.2% of the level-2 random intercept variance and 0% of the level-1 residual print("Rights & Sterba R2 suite") variance (as expected for school-level predictors). r2mlm(model=Model2, bargraph=FALSE) \$R2s total within between fl 0.00000000 **0** NA  $\rightarrow$  Pseudo-R2 for level-1 residual NA 0.24306893 > Pseudo-R2 for level-2 random intercept f2 0.044375381 0.000000000 0 NA V 0.138187567 NA 0.75693107 m **Total-R<sup>2</sup> Results:** The 2 fixed slopes of our 2 level-2 predictors f 0.044375381 NA NA accounted for 4.33% of the total math variance (approximated by fv 0.044375381 0 NA  $.232^*$ . 180 = 4.2% using the ICC from the empty model). fvm 0.182562948 NA NA print("Total-R2 relative to empty model using Jonathan's function") totalRSquaredinator(model=Model2, dvName="math", data=Example2) 0.043102056 # Compute total-R2 the longer way instead by saving predicted outcomes Example2\$PredMain = predict(Model2, re.form=NA) # Do not include random intercept rModel2 = cor.test(Example2\$PredMain, Example2\$math, method="pearson")

print("Total R2"); rModel2\$estimate^2 0.043102056

# Model 3: Add Interaction of Level-2 School Size and Public vs Private (6 parameters)

Level 1:  $Math_{pc} = \beta_{0c} + e_{pc}$ Level 2:  $\beta_{0c} = \gamma_{00} + \gamma_{01}(size_c - 100) + \gamma_{02}(private_c) + \gamma_{03}(size_c - 100)(private_c) + U_{0c}$ Slope of school size for private schools =  $[\gamma_{01} + \gamma_{03}(private_c)](size_c - 100)$ 

# **STATA** Syntax and Output for Model 3:

#### 

The ## in the fixed effects means "give me the two-way interaction and all lower-order main effects" (for less typing).

Log restricted-like	elihood = -2	3541.699 <b>= L</b>	F( L Pr	<b>3, 15</b> ob > F	4.92) = =	<b>15.26</b> 0.0000	
math	n   Coe	f. Std. Er	r.	DF	t	P> t	
size100 private c.size100#c.private 	)   .03066   3.2518   .12393   11.287	22 .044786 38 .493514 53 .090033 55 .329090	2 8 6 6	160.8 153.0 152.3 158.4	0.68 6.59 1.38 34.30	0.495 gam 0.000 gam 0.171 gam 0.000 gam	ma01 ma02 ma03 ma00
Random-effects Pa	arameters	Estimate	 Std. E		[95% Conf.	Interval]	
schoolID: Identity	//////////////////////////////////////	6.569508	.85824	94	5.085466	8.486624	Var(U_0c)
var	(Residual)	39.1507	.66072	27	37.87689	40.46734	Var(e_pc)
estat ic, n(160) Akaike's informatio Model	on criterion  Obs ll(	// AIC and and Bayesia null) ll(mo	BIC for n inform 	ation c df	<b>1-2 units</b> riterion AIC	BIC	
Model	Obs 11(	null) ll(mo	 del) 	df	AIC	BIC	
.	160	235	41.7	6	47095.4	47113.85	
lincom c.size100*1	+ c.size100	#c.private /	/ Size s	lope fo	r private s	chools	
math	Coef. St	d. Err.	z P>	z	[95% Conf.	Interval]	
(1)   .2	.0	781041 1	.98 0.	048	.0015163	.3076787	g01 + g03(1)
predict predinterad corr math predinter	et cact math predin	// Save fi // Get tot ~t	xed-effe al r to	ect pred make R2	icted outco	mes	
math   1.	.0000 .2130 1.00	00					

display "Total-R2 = " r(rho)^2 // Print total R2 relative to empty model Total-R2 = .04537771

#### **R** Syntax and Output for Model 3:

```
print("R Model 3: Add Interaction of Level-2 School Size and Public vs Private")
Model3 = lmer(data=Example2, REML=TRUE, formula=math~1+size100*private+(1|schoolID))
print("Show results using Satterthwaite DDF including -2LL as deviance")
llikAIC(Model3, chkREML=FALSE); summary(Model3, ddf="Satterthwaite")
```

The \* in the fixed effects means "give me the two-way interaction and all lower-order main effects" (for less typing).

```
'log Lik.' -23541.699 (df=6) → LL for model (with 6 parameters)
$AICtab
                  BIC
                           loqLik
                                    deviance
                                               df.resid
       AIC
 47095.398 47136.677 -23541.699 47083.398 7179.000 → deviance = -2LL for model
Random effects:
 Groups Name
                     Variance Std.Dev.
 schoolID (Intercept) 6.5695 2.5631 Var(U_0c)
Residual
                      39.1507 6.2571 Var(e pc)
Number of obs: 7185, groups: schoolID, 160
Fixed effects:
                  Estimate Std. Error
                                               df t value
                                                                  Pr(>|t|)
                 11.287555 0.329090 155.263861 34.2993
(Intercept)
                                                                 < 2.2e-16 gamma00
                  0.030662 0.044786 157.645142 0.6846
size100
                                                                    0.4946 gamma01
                  3.251838 0.493513 150.004059 6.5892 0.000000007046 gamma02
private
size100:private 0.123935 0.090033 149.283221 1.3765
                                                                    0.1707 gamma03
Intercept \gamma_{00} =
size100 \gamma_{01} =
private \gamma_{02} =
size100*private \gamma_{03} =
print("Size slope for private schools?"); contest1D(Model3, L=c(0,1,0,1))
    Estimate Std. Error
                                df t value
                                                 Pr(>|t|)
1 0.15459752 0.078103868 146.67772 1.9793837 0.049644738 gamma01 + gamma03(1)
print("F-Test of Model Total-R2")
contestMD(Model3, ddf="Satterthwaite", L=rbind(c(0,1,0,0),c(0,0,1,0),c(0,0,0,1)))
             Mean Sq NumDF DenDF F value
     Sum Sq
                                                              Pr(>F)
1 1792.4576 597.48586
                        3 151.89378 15.261179 0.0000000099242283
print("Psuedo-R2 relative to empty model using Jonathan's function")
pseudoRSquaredinator(smallerModel=Model1, largerModel=Model3)
R2 Random.(Intercept) R2 L1.sigma2
       0.237352159809
                             -0.000060781529
print("Rights & Sterba R2 suite -- f2 is for random intercept")
r2mlm(model=Model3, bargraph=FALSE)
$R2s
                                        Pseudo-R<sup>2</sup> Results: The 3 fixed slopes of our 3 level-2
          total within
                          between
                                        predictors accounted for 23.7% of the level-2 random
f1 0.00000000
                    0
                               NA
                    NA 0.25496998
                                        intercept variance, an increase of 0.5%, using the usual
f2 0.046869466
                                        way of computing it. The alternative (from just this
    0.00000000
                    0
                               NA
v
   0.136954000
                    NA 0.74503002
                                        model using r2mlm) provides a slightly higher estimate.
m
f
    0.046869466
                    NA
                               NA
fv 0.046869466
                    0
                               NA
fvm 0.183823466
                    NA
                               NA
print("Total-R2 relative to empty model using Jonathan's function")
```

totalRSquaredinator(model=Model3, dvName="math", data=Example2)

0.04537771

**Total-R<sup>2</sup> Results:** The 3 fixed slopes of our 3 level-2 predictors accounted for 4.53% of the total math variance, as increase of 0.2% relative to the main effects model.

# Sample Results Section [indicates notes about what to customize or also include; note that SE and *p*-values are not needed if you provide a table for the model solution]

The extent to which the extent to which student math outcomes (M = 12.75, SD = 6.88, range = -2.83 to 24.99) could be predicted from school-level variables of school size and public versus private status was examined in a series of multilevel models in which the 7,185 students were modeled as nested within their 160 schools. The number of students included per school ranged from 14 to 67 (M = 45). Restricted Maximum likelihood (REML) within SAS MIXED [or STATA MIXED or R lmer] was used in estimating and reporting all model parameters. The significance of fixed effects was evaluated with univariate and multivariate Wald tests using Satterthwaite denominator degrees of freedom, whereas random effects were evaluated via likelihood ratio tests (i.e.,  $-2\Delta LL$  with degrees of freedom equal to the number of new random effects variances and covariances). Alpha was chosen as .01. Model-implied fixed effects were requested via ESTIMATE [or LINCOM or contest1D] statements. Effect size for the fixed effects was evaluated via pseduo-R<sup>2</sup> values for the proportion reduction in each variance component relative to a nested model without the predictors in question, as well as with total-R<sup>2</sup>, the squared correlation between the actual math outcomes and those predicted by the model fixed effects.

As derived from an empty means, random intercept model, student math had an intraclass correlation of ICC = .180, indicating that 18.0% of the variance in student math was between schools, a significant amount,  $-2\Delta LL(1) = 986.12$ , p < .0001. Given an average of 45 students per school in this sample, the ICC = .180 translated into a design effect = 7.36, further indicating the need for a multilevel analysis. The school mean math outcomes had strong reliability, as evidenced by a weighted ICC2 = .908. The fixed intercept was 12.637 (SE = 0.244), which represented the expected average school mean math outcome. A random intercept confidence interval (computed as the fixed intercept  $\pm$  1.96\*SQRT[random intercept variance]) indicated that 95% of the schools were expected to have school mean math outcomes between 6.884 and 18.390 (around the average of 12.637).

We then added the school-level predictors of school size (per 100 students, centered so that 0 = 1000 students) and status (public = 0 versus private = 1). These two fixed effects accounted for significant variance overall, F(2, 157.31) = 21.80, p < .0001, including 23.2% of the level-2 random intercept variance and 4.3% of the total variance. The fixed intercept was 11.182 (SE = 0.321), which represented the expected school mean math outcome for public schools with 1000 students. The slope for school size was nonsignificantly positive, indicating that school mean math was expected to be nonsignificantly higher by 0.061 (SE = 0.039, p = .118) per 100 more students. The slope for public vs. private was significantly positive, indicating school mean math was predicted to be higher by 3.151 (SE = 0.490, p < .0001) for private schools relative to public schools.

We then added an interaction between school size and type. The three fixed effects still accounted for significant variance overall, F(3, 154.92) = 15.26, p < .0001, including 23.7% of the level-2 random intercept variance (an increase of 0.5%) and 4.5% of the total variance (an increase of 0.2%). The interaction slope was nonsignificantly positive, indicating that the slope of school size for private schools (Est = 0.155, SE = 0.078, p = .495) was 0.12 nonsignificantly more positive than the size slope for public schools (Est = 0.031, SE = 0.045, p = .050). The slope for public vs. private remained significantly positive, indicating school mean math was predicted to be higher by 3.252 (SE = 0.494, p < .0001) for private schools relative to public schools.